#### Discrimination between Gaussian process models: active learning and static constructions

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# Introduction

Investigating the properties of static as well as incremental/sequential design criteria for discriminating between the correlation structure of (two) Gaussian process models.

- T-optimality (Atkinson and Fedorov, 1975) not applicable since it assumes iid normal data with constant variance.
- Instead, one may use the symmetrized Kullback-Leibler (KL) divergence between the two models as criterion.
- Symmetrized KL divergence computationally expensive if many design points ⇒ develop new criteria inspired by Fréchet distance.
- ► For these new design criteria it is straightforward to introduce design measures and derive necessary conditions for optimality ⇒ possible to apply approximate design methods.

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Gaussian processes			

### Gaussian processes/fields

- Marginals of Gaussian fields at any subset of points/locations have a multivariate normal distribution.
- Frequently used e.g. as surrogate models for computer experiments (Gramacy, 2020) or in machine learning.
- Gaussian process regression / kriging (Stein, 1999): assume Gaussian process prior and obtain distribution for "true" function at unseen locations given observed points.

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Gaussian processes			

## Gaussian processes/fields

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- General setting and notation:
  - ▶ random field  $Z_x$ , indexed by  $x \in \mathscr{X} \subset \mathbb{R}^d$ .
  - Y(x): realization of the random field.
  - ►  $E{Z_x} = 0 \forall x \text{ and } E{Z_xZ'_x} = K(x,x') \forall (x,x') \in \mathscr{X}^2.$
  - kernel K(x, x') = σ<sup>2</sup> f<sub>θ</sub> (||x − x'||): isotropic, continuous and decreasing function of the distance.

We do not consider repeated observations (no nugget effect).

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## Symmetrized Kullback-Leibler divergence

Given two probability density functions φ<sub>0</sub>(y, θ<sub>0</sub>) and φ<sub>1</sub>(y, θ<sub>1</sub>), maximise the expected power of the likelihood ratio test if model 1 is the true model (see, e.g., López-Fidalgo et al., 2007):

$$\mathsf{E}_1(L) = \int \varphi_1(y, \theta_1) \log \left\{ \frac{\varphi_1(y, \theta_1)}{\varphi_0(y, \theta_0)} \right\} \mathrm{d}y = D_{\mathsf{KL}}(\varphi_1 \| \varphi_0)$$

Now maximise the power of the likelihood ratio test if model 0 is the true model:

$$\mathsf{E}_{0}(-L) = \int \varphi_{0}(y,\theta_{0}) \log \left\{ \frac{\varphi_{0}(y,\theta_{0})}{\varphi_{1}(y,\theta_{1})} \right\} \mathrm{d}y = D_{\mathsf{KL}}(\varphi_{0} \| \varphi_{1})$$

The symmetrized KL divergence is the average of these two divergences (see, e.g., Pronzato et al., 2019):

$$D_{\mathcal{KL}}(\varphi_0,\varphi_1) = \frac{1}{2} \left[ D_{\mathcal{KL}}(\varphi_0 \| \varphi_1) + D_{\mathcal{KL}}(\varphi_1 \| \varphi_0) \right]$$

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## Symmetrized KL divergence for Gaussian random field

- The two models differ through their kernel functions.
- Given the *n*-point design X<sub>n</sub> = (x<sub>1</sub>,..., x<sub>n</sub>), construct the kernel matrix for model *i* (*i* = 0, 1), K<sub>n,i</sub>, as {K<sub>n,i</sub>}<sub>j,k</sub> = K(x<sub>j</sub>, x<sub>k</sub>) for 1 ≤ j, k ≤ n.
   For the Gaussian random field.

$$\varphi_{n,i}(\mathbf{Y}_n) = \frac{1}{(2\pi)^{n/2} \det^{1/2} \mathbf{K}_{n,i}} \exp\left[-\frac{1}{2} \mathbf{Y}_n^\top \mathbf{K}_{n,i}^{-1} \mathbf{Y}_n\right], \ i = 0, 1.$$

Therefore,

$$\Phi_{\mathcal{K}L[\mathcal{K}_{0},\mathcal{K}_{1}]}(\mathbf{X}_{n}) = 2 D_{\mathcal{K}L}(\varphi_{n,0},\varphi_{n,1}) = \\ = \frac{1}{2} \left[ \operatorname{trace}(\mathbf{K}_{n,0}\mathbf{K}_{n,1}^{-1}) + \operatorname{trace}(\mathbf{K}_{n,1}\mathbf{K}_{n,0}^{-1}) \right] - n.$$

 Disadvantages: cumbersome and unstable computation (matrix inverses), no generalisation to design measures.

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## Fréchet distance and $\Phi_p$ criteria

 Consider alternatively the Fréchet distance, related to the Wasserstein distance (Dowson and Landau, 1982):

$$\Phi_{F[\mathcal{K}_0,\mathcal{K}_1]}(\mathbf{X}_n) = \operatorname{trace}\left[\mathbf{K}_0 + \mathbf{K}_1 - 2\left(\mathbf{K}_0\mathbf{K}_1\right)^{1/2}\right].$$

This is also difficult to compute, but squaring all matrices gives

$$\Phi_{2[\mathcal{K}_{0},\mathcal{K}_{1}]}(\boldsymbol{X}_{n}) = \operatorname{trace}\left(\boldsymbol{K}_{0}^{2} + \boldsymbol{K}_{1}^{2} - 2\,\boldsymbol{K}_{0}\boldsymbol{K}_{1}\right) = \operatorname{trace}\left[(\boldsymbol{K}_{0} - \boldsymbol{K}_{1})^{2}\right].$$

• *Idea*: introduce criteria  $\Phi_{p[K_0,K_1]}$  defined as

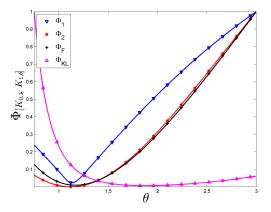
$$\Phi_{p[\mathcal{K}_{0},\mathcal{K}_{1}]}(\mathbf{X}_{n}) = \|\mathbf{K}_{1} - \mathbf{K}_{0}\|_{p}^{p} = \sum_{i,j=1}^{n} |\{\mathbf{K}_{1} - \mathbf{K}_{0}\}_{i,j}|^{p}, \ p > 0.$$

### Example setup

- $\mathscr{X} = [0, 10]^2$ , grid size:  $25 \times 25$ .
- Rival models: Matérn family ( $\sigma^2 = 1$ ):
  - model 0: Matérn 3/2
  - model 1: Matérn 5/2
- We consider discrimination designs only for fixed parameters (locally optimum designs).
- Find inverse length-scales  $\theta_0$  and  $\theta_1$  where both models agree most:
  - Take θ<sub>0</sub> = 1 in the first kernel, adjust the parameter in the second kernel minimizing Φ<sub>1</sub>, Φ<sub>2</sub>, Φ<sub>F</sub>, and Φ<sub>KL</sub> for the design X<sub>625</sub>.
  - Results:  $\theta_1 = 1.0047, 1.0285, 1.0955, 1.3403$ , resp.
  - We have finally used the setting  $\theta = (1, 1.07)$ .

Criteria for exact designs

## Selection of parameters

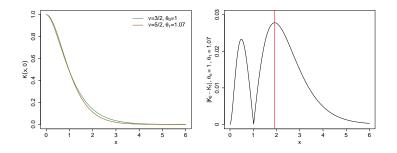


 $\Phi_1$ ,  $\Phi_2$ ,  $\Phi_F$ ,  $\Phi_{KL}$  as functions of  $\theta$  for an equally-spaced 11-point design.

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Criteria for exact designs

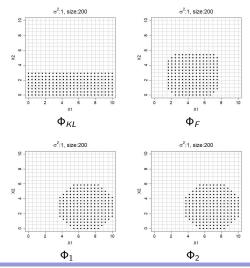
## Covariance functions at selected parameters



Left: Plot of the Matérn covariance functions. Right: Absolute difference of covariance functions at distance *x*.

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### Exact designs for static criteria



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#### Performance of criteria in simulation study

- For each design of size n, N = 100 independent sets of n observations are simulated from the "true" model 0.
- The hit rate is computed as the proportion of samples where the likelihood of model 0 is larger than the likelihood of model 1.
- The same is repeated for the case where model 1 is the correct model.
- The two hit rates are averaged to obtain the average hit rates.

	Average hit rate									
Design size	5	6	7	8	9	10	20	30	40	50
Φ <sub>F</sub>	0.580	0.625	0.620	0.625	0.670	0.715	0.795	0.900	0.925	0.950
$\Phi_1$	0.525	0.520	0.555	0.540	0.550	0.610	0.725	0.890	0.910	0.920
Φ <sub>2</sub>	0.525	0.520	0.555	0.540	0.550	0.610	0.715	0.860	0.890	0.910
Φ <sub>KL</sub>	0.580	0.625	0.620	0.625	0.670	0.715	0.795	0.895	0.925	0.955

Criteria for approximate designs

## Design measure version of $\Phi_p$ criterion

• Defining  $\xi_n$  as the empirical measure on the points in  $\mathbf{X}_n$ ,  $\xi_n = (1/n) \sum_{i=1}^n \delta_{x_i}$ , one can write

$$\Phi_{\rho[\mathcal{K}_0,\mathcal{K}_1]}(\mathbf{X}_n) = n^2 \phi_{\rho[\mathcal{K}_0,\mathcal{K}_1]}(\xi_n),$$

where

$$\phi_{p[K_0,K_1]}(\xi) = \int_{\mathscr{X}^2} |K_1(x,x') - K_0(x,x')|^p \,\mathrm{d}\xi(x) \mathrm{d}\xi(x').$$

Criteria for approximate designs

# Necessary condition for optimality

#### Theorem

If the probability measure  $\xi^*$  on  $\mathscr{X}$  maximises  $\phi_{p[K_0,K_1]}(\xi)$ , then

$$\forall x \in \mathscr{X}, \ \int_{\mathscr{X}} |\mathcal{K}_1(x, x') - \mathcal{K}_0(x, x')|^p \, \mathrm{d}\xi^*(x') \leq \phi_{p[\mathcal{K}_0, \mathcal{K}_1]}(\xi^*).$$

Moreover,  $\int_{\mathscr{X}} |K_1(x, x') - K_0(x, x')|^p d\xi^*(x') = \phi_{p[K_0, K_1]}(\xi^*)$  for  $\xi^*$ -almost every  $x \in \mathscr{X}$ .

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Simplified problem with explicit solution for optimum

• Let 
$$K_i(x, x') = \Psi_i(||x - x'||)$$
,  $i = 0, 1$ , and  $\psi(t) = |\Psi_1(t) - \Psi_0(t)|$ ,  $t \in \mathbb{R}^+$ .

• Then 
$$\phi_p(\xi) = \int_{\mathscr{X}^2} \psi(\|x - x'\|)^p d\xi(x) d\xi(x').$$

• Consider the extreme case  $\psi = \psi_*$  defined by

$$\psi_*(t) = \left\{egin{array}{cc} 1 & ext{if } t = \Delta, \ 0 & ext{otherwise.} \end{array}
ight.$$

Since  $\psi_*(t)^p = \psi(t)^p$  for any p > 0, we only need to consider p = 1.

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#### Theorem

When  $\psi = \psi_*$  and  $\mathscr{X} \subset \mathbb{R}^d$  is large enough to contain a regular d simplex with edge length  $\Delta$ , any measure  $\xi^*$  allocating weight 1/(d+1) at each vertex of such a simplex maximises  $\phi_1(\xi)$ , and  $\phi_1(\xi^*) = d/(d+1)$ .

Criteria for approximate designs

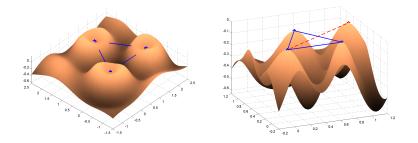
Are results for  $\psi = \psi_*$  generalisable?

- The results for ψ = ψ<sub>\*</sub> do not generalise to the general function ψ(t) = |Ψ<sub>1</sub>(t) − Ψ<sub>0</sub>(t)|.
- Even if p → ∞, one can show that one can always find a better design than ξ\*, given the design space is large enough.
- However, the simplex design might be close to optimal (at least for high p).

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Criteria for approximate designs

## Illustration of directional derivative



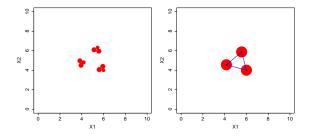
Surface plots of directional derivatives: Left:  $K_0 = K_{0,1}$ ,  $K_1 = K_{1,1.07}$  ( $\Delta \simeq 1.92$ ), p = 2. Right:  $K_0 = K_{0,1}$ ,  $K_1 = K_{1,1}$  ( $\Delta \simeq 0.7$ ), p = 10.

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## Numerical optimisation

Two-step approach:

- 1. Run Fedorov-Wynn algorithm (Fedorov, 1971; Wynn, 1970) for 1000 iterations using directional derivatives from necessary condition on a dense regular grid.
- 2. Run continuous optimisation algorithm for coordinates and design weights starting from design found in step 1.



Left: The optimal measure for  $\phi_2$ . Right: The optimal measure for  $\phi_{10}$ .  $\theta_1 = 1.07$ . The edge lengths of the triangles are  $\Delta \simeq 1.92$ .

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## Sequential/incremental designs

We also compared the static criteria to the following construction methods:

Sequential design:

put next observation(s) where symmetrized KL divergence between predictive distributions of the models differs most.

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Sequential design:

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- Incremental design:
  - Incrementally build design by putting next point where (normalised) differences between the prediction errors of the two models are largest.
  - Alternatively use symmetrized KL divergence (cond. on current design).
  - Theoretical investigations:
    - prediction-based and KL criteria tend to behave differently and depend on covering radius (CR) relative to correlation length.
    - KL divergence sometimes clearly better suited for discrimination (remains positive with decreasing CR).

# Conclusion

- Introducing a new family of criteria which are simple to compute and allow for a formulation in terms of approximate design measures.
- In our examples, they lead to marginally worse performance than symmetrized KL divergence.
- For large p, designs with d + 1 support points placed on the vertices of a simplex are often (close to) optimal for the new criteria, depending on the size of the design space.
- Not considered yet: behaviour under parameter uncertainty.

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